

## Parameter estimation L Optimization

- Estimation of parameters appearing in models / functions
- Efficient use of data for mathematical modelling
- simplest model linear regression

• Let  $\vec{p}$  be the parameter vector containing the ~~form~~ of parameters to be optimized

$m$ : is number of parameters  $p_1, \dots, p_m$

• Let  $\vec{z}$  be the ~~measured~~ vector which is the "ideal" output of the system to be modeled

the system in the noise-free case is described by a vector function  $f$  which relates  $\vec{z}$  to  $\vec{p}$  and that

$$f(\vec{p}, \vec{z}) = 0$$

• In practice measurements  $\vec{y}$  are only available for system output  $\vec{z}$  with noise

$$\vec{y} = \vec{z} + \vec{\varepsilon}$$

• We take multiple measurements of the system  $\{y_i\}_{i=1, \dots, n}$  and want to estimate  $\vec{p}$  using  $\{y_i\}$   
→ due to noise  $f(\vec{p}, y_i) = 0$  is not valid anymore

~~the~~ Solution: We write cost function or objective function  $F$  describing the error between measurement and system output for given parameters

$$F(\vec{p}, y_1, \dots, y_n)$$

and minimize the cost

If there are no constraints on  $\vec{p}$  and function  $F$  has first and second order partial derivatives, necessary conditions for a minimum are

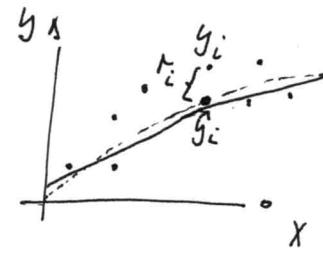
$$\frac{\partial F}{\partial \vec{p}} = 0 \quad \text{and} \quad \frac{\partial^2 F}{\partial \vec{p}^2} > 0$$

## Least-squares optimization

Minimize the error of sum of squares

$$\min_{\vec{p}} F = \sum_{i=1}^n (y_i - \hat{y}_i)^2 \quad \text{with } \hat{y}_i = f(x_i; \vec{p}) \text{ model prediction}$$

$$F(\vec{p}) = \sum_{i=1}^n (y_i - f(x_i; \vec{p}))^2$$



$$y_i = p_0 + p_1 x_i + \varepsilon_i \quad \varepsilon_i \sim N(0, \sigma^2)$$

$p_0, p_1$  so werten dass  $y(x) = \hat{p}_0 + \hat{p}_1 x$   
möglichst nah an den Daten liegt  
d.h. Abstand minimieren

if the errors are normally distributed

the least square estimator are also the maximum likelihood estimator

How to find  $\hat{p}$  which minimizes  $F$ :

1. Finding an analytical solution by

- differentiating  $F$  with respect to  $p_1, \dots, p_m$
- setting partial derivatives zero  $\frac{\partial F}{\partial p_i} = 0$
- solving the resulting  $m$  normal equations
- only working for very few nonlinear models

## 2 Numerical Solutions

- try different values for parameters  $\hat{p}_g$
- calculate  $F(\hat{p}_g)$  and work towards smaller  $F(\hat{p}_g)$

3 main procedures (sensitivity based approach)

Steepest descent method (gradient descent)

Searches minimum  $F$  by iteratively determining the direction in which the parameter estimates should change

$F(\hat{p})$  is defined and differentiable

$F(\hat{p})$  decreases fastest in point  $\hat{p} = \hat{a}$  if one goes from  $\hat{a}$  in the direction of negative gradient of  $F$  at  $\hat{a}$

$$-\nabla F(\hat{a})$$

$$\hat{a}_{k+1} = \hat{a}_k - \gamma \nabla F(\hat{a}_k) \quad \begin{matrix} \text{if } \gamma \text{ the step size} \\ \text{allowed to change} \end{matrix}$$

Gauss-Newton method (GNA) or linearization

- uses Taylor series expansion to approximate the nonlinear model with linear terms.

Terms are used in least regression to come up with new terms

(Levenberg) Marquardt algorithm or damped least squares

(interpolator between Gauss-Newton & steepest descent algorithm)

Starting values:

- all iterative procedures require starting values
- risk of local minima (multiple start methods, particle swarm)



Important

- The simpler the model, the better the behavior in the estimation process.
- Over parameterization often leads to convergence problems
  - .. may have multiple solutions
  - .. high correlation between parameter estimates

Parameter can have large influences

$$y_i = \frac{\beta_0 x_i}{\beta_1 x_i + \beta_2} \neq y_i \quad \text{vs.} \quad y_i = \frac{x_i}{c_0 x_i + c_1}$$

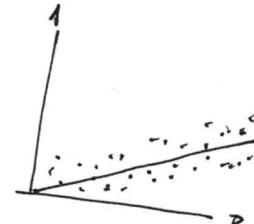
Practical example:

linear regression

$$y_i = \hat{p}_1 x_i + \varepsilon_i \quad \varepsilon_i \sim N(0, \sigma^2)$$

$$F(p_1) = \sum_{i=1}^n (y_i - \hat{y}_i)^2$$

$$\min \sum_{i=1}^n (y_i - p_1 x_i)^2$$



$$\frac{\partial F}{\partial p_1} = \sum_{i=1}^n \cancel{2} (y_i - p_1 x_i) \cdot (-x_i) * *$$

$$\sum_{i=1}^n (-2 x_i y_i + 2 \cancel{p_1 x_i^2}) = 0$$

$$\cancel{\sum x_i y_i} = \cancel{\sum p_1 x_i^2} = \cancel{p_1 \sum x_i^2}$$

$$\frac{\cancel{\sum x_i y_i}}{\sum x_i^2}$$

$$\frac{\partial^2 F}{\partial^2 p_1} = \sum_{i=1}^n x_i^2 \rightarrow 0$$

$$\begin{aligned} p_0 &= \langle \hat{y}_i - b_1 \langle x \rangle, \\ p_1 &= \frac{\sum_{i=1}^n (x_i - \langle x \rangle)(y_i - \langle y \rangle)}{\sum_{i=1}^n (x_i - \langle x \rangle)^2} \\ &= \frac{\text{Cov}(X, Y)}{\text{Var}(X)} \end{aligned}$$

From where come the ~~derivative~~ derivatives from (for sensitivity based methods)

$$\frac{\partial F}{\partial p_i}$$

- often finite differences, but very numerical intensive

$$\frac{\partial F}{\partial p_i}(\vec{a}) \approx \frac{F(\vec{a}) - F(\vec{a} + \Delta p_i)}{P_i - \Delta p_i}$$

- based on sensitivity equations  
analytical; e.g. use ODE's +  $F$  and perform derivatives (Symbolic math)

### Stochastic approaches:

- simulated annealing
- particle swarm - population of candidate solutions

### Bayesian approaches: posterior distributions

maximum likelihood

+ Gibbs sampling / MCMC sampling

### Identifiability: profile likelihood

### Local vs global maximum: Waterfall plots